### New Modeling Methods for Multilevel Data

#### Tihomir Asparouhov and Bengt Muthén

May 23, 2011

Tihomir Asparouhov and Bengt Muthén Muthén & Muthén 1/57

- Two-level modeling for categorical data: weighted least squares estimation
- Three-level structural equation modeling: maximum-likelihood estimation
- Bayesian estimation for two and three level modeling
- Three-level modeling for categorical data: Bayesian estimation
- Cross classified structural equation models
- Two-level missing data imputation
- Two-level exploratory factor analysis
- Two-level multiple group modeling
- Plausible values for random effects

### Two-level modeling for categorical data: weighted least squares estimation

- Asparouhov, T. & Muthén, B. (2007). Computationally efficient estimation of multilevel high-dimensional latent variable models. Proceedings of the 2007 JSM.
- Extends the single level methodology originated by Muthén (1978) Contributions to factor analysis of dichotomous variables.
- Mplus Version 5 2007
- Advantages over ML: unlimited number of latent variables, fast computation, useful chi-square test of fit
- Disadvantages over ML: models only continuous and categorical variables, MAR missing data not supported, can yield non-positive definite variance/covariance for large number of random effects, only random intercepts.

### Two-level WLS: Model

$$y_{pij} = k \Leftrightarrow \tau_{pk} < y_{pij}^* < \tau_{pk+1}.$$
 (1)

$$y_{pij}^* = y_{wpij} + y_{bpj} \tag{2}$$

$$y_{wij} = \Lambda_w \eta_{wij} + \varepsilon_{wij} \tag{3}$$

$$\eta_{wij} = B_w \eta_{wij} + \Gamma_w x_{wij} + \xi_{wij}.$$
(4)

$$y_{bj} = \mathbf{v}_b + \Lambda_b \eta_{bj} + \varepsilon_{bj} \tag{5}$$

$$\eta_{bj} = \alpha_b + B_b \eta_{bj} + \Gamma_b x_{bj} + \xi_{bj}.$$
 (6)

#### Part 1. Estimate the unrestricted model

$$y_{pij} = k \Leftrightarrow t_{pk} < y_{pij}^* < t_{pk+1}.$$
(7)

$$y_{pij}^* = y_{wpij} + y_{bpj}.$$
 (8)

$$y_{wij} = \beta_w x_{wij} + \varepsilon_{wij} \tag{9}$$

$$y_{bj} = \mu_b + \beta_b x_{bj} + \varepsilon_{bj}.$$
 (10)

- ML Univariate estimation 1 dimension of numerical integration
- ML Bivariate estimation 2 dimensions of numerical integration estimation only for 2 correlation parameters

Part 2. Estimate the structural model by minimizing

$$F = (s - s^*)W(s - s^*)^T$$
(11)

- s are the parameters of the unrestricted model
- $s^*$  are the implied quantities from the structural model
- W is a weight matrix

- F is the base of the chi-square test of fit for the structural model v.s. the unrestricted model.

- 6 polytomous observed variables with 5 categories
- 100 clusters of size 10
- 2 within and 2 between factors in a two-level CFA each measured by 3 variables
- ML uses 8 dimensional integration montecarlo integration

### Two-level WLS: Simulation Study Results

#### Table: Two-level factor analysis model with categorical variables.

para-	true	WLSM	ML	WLSM	ML	Efficiency
meter	value	bias	bias	coverage	coverage	ratio
$\lambda_{w2}$	1.0	3%	2%	97%	100%	1.14
$\psi_{w12}$	0.4	2%	-14%	97%	89%	0.89
$\psi_{w11}$	0.7	2%	-23%	94%	75%	0.71
$\lambda_{b2}$	1.0	5%	4%	96%	94%	0.96
$\psi_{b12}$	0.2	-1%	-22%	94%	81%	0.91
$\psi_{b11}$	0.4	1%	-31%	93%	57%	0.77
$ au_{11}$	-0.3	-3%	-6%	96%	87%	1.17
$ au_{12}$	0.4	-1%	-14%	96%	81%	1.00
$ au_{13}$	1.2	0%	-11%	95%	55%	0.71
$ au_{14}$	1.8	0%	-10%	98%	47%	0.56
$\theta_{b1}$	0.2	-2%	-55%	97%	32%	0.66

### Three-level structural equation modeling: maximum-likelihood estimation

### 3-level SEM

- New feature in Mplus 7
- Three level multivariate data  $Y_{pijk}$

$$Y_{pijk} = Y_{1pijk} + Y_{2pjk} + Y_{3pk}$$
(12)

• 3 sets of structural equations - one on each level

$$Y_{1ijk} = \Lambda_1 \eta_{ijk} + \varepsilon_{ijk} \tag{13}$$

$$\eta_{ijk} = B_1 \eta_{ijk} + \Gamma_1 x_{ijk} + \xi_{ijk}.$$
<sup>(14)</sup>

$$Y_{2jk} = \Lambda_2 \eta_{jk} + \varepsilon_{jk} \tag{15}$$

$$\eta_{jk} = B_2 \eta_{jk} + \Gamma_2 x_{jk} + \xi_{jk}. \tag{16}$$

$$Y_k = \mathbf{v} + \Lambda_3 \eta_k + \varepsilon_k \tag{17}$$

$$\eta_k = \alpha + B_3 \eta_k + \Gamma_3 x_k + \xi_k. \tag{18}$$

### **3-level SEM: Estimation**

- Double EM-algorithm where the latent variables are  $Y_{2jk}$  and  $Y_{3k}$
- E-step is based on  $[Y_{3k}|*]$  and  $[Y_{2jk}|Y_{3k},*]$
- M-step is a simple 3 group SEM maximization.
- Missing data: MAR support. Not a part of the EM-algorithm.
- Model accommodates easily variables defined on different levels: 7 types.
- Random coefficients for level 1 covariates and level 2 covariates

$$\Gamma_1 = \gamma_1 + \gamma_{1jk} + \gamma_{1k}$$

$$\Gamma_2 = \gamma_2 + \gamma_{2k}$$

where  $\gamma_{1jk}$  is a level 2 latent variable while  $\gamma_{1k}$  and  $\gamma_{2k}$  are level 3 latent variables.

- No numerical integration
- Fast computation all runs less than 1 min.
- Robust standard errors using sandwich estimator.

### 3-level SEM Example 1: Path Analysis / 3-level regression

$$\begin{aligned} Y_{ijk} &= Y_{1ijk} + Y_{2jk} + Y_{3k} \\ Z_{ijk} &= Z_{1ijk} + Z_{2jk} + Z_{3k} \\ Y_{1ijk} &= \beta_1 Z_{1ijk} + \varepsilon_{ijk} \\ Y_{2jk} &= \beta_2 Z_{2jk} + \varepsilon_{jk} \\ Y_{3k} &= \alpha + \beta_3 Z_{3k} + \varepsilon_k \end{aligned}$$

- Simulation: 80 level 3 clusters, with 15 level 2 clusters each of size 10.
- Extension of BFSPE, bias reduction (observed v.s. latent predictor), and Mediation modeling to 3 level
- Marsh et al. (2009). Doubly-latent models of school contextual effects: Integrating multilevel and structural equation approaches to control measurement and sampling errors. MBR, 44, 764-802.
- Preacher et al. (2010). A general multilevel SEM framework for assessing multilevel mediation. Psychological Methods, 15,

209-233.

#### 3-level SEM Example 1: Path Analysis Results

MODEL RESULTS

	Des	oulation	ESTIMATES	Std. Dev.	S. E. Average	M. S. E.	95% Cover	
	thin Leve		Average	Sta. Dev.	Average		Cover	Co
	uru reve	-						
¥1	ON							
¥2		0.400	0.3987	0.0101	0.0094	0.0001	0.920	1.
Varianc	es							
¥2		1.200	1.1991	0.0152	0.0162	0.0002	0.960	1.
Residua	1 Varian	ces						
¥1		1.200	1.1999	0.0136	0.0160	0.0002	0.980	1.
• Bet	ween CLU	STER1 Lev	el					
¥1	ON							
¥2		0.300	0.2933	0.0362	0.0352	0.0013	0.930	1.
Varianc	es							
¥2		0.700	0.7013	0.0362	0.0342	0.0013	0.940	1.
Residua	1 Varian							
¥1		0.700	0.7062	0.0375	0.0343	0.0014	0.920	1.
• Bet	ween CLU	STER2 Lev	el					
¥1	ON							
¥2		0.100	0.0692	0.1343	0.1271	0.0188	0.900	0.
Means								
¥2		0.000	-0.0073	0.0736	0.0746	0.0054	0.950	0.
Interce	pts							
¥1		0.000	-0.0094	0.0689	0.0756	0.0048	0.960	0.
Varianc	es							
¥2		0.400	0.3938	0.0707	0.0703	0.0050	0.920	1.
Residua	l Varian	ces						
¥1		0.400	0.3920	0.0653	0.0680	0.0043	0.920	1.
		1.22	engt Muthé		thén & Mu	a 2 -	14/ 57	

- Simulation: 1 factor on each level with 5 indicators.
- 3 simulations with different number of level 3 clusters: 80, 40, 20.
- Chi-square test of fit between the restricted model factor analysis model and the unrestricted variance covariance.
- DF=15, average chi-square value 15.156 / 15.571 / 17.707, rejection rates 5% / 9% / 19%

# 3-level SEM Example 2: Factor Analysis Results, 80 level 3 clusters

MODEL RESULTS

Po Within Leve E BY Y2 Variances E Residual Varian Y1 • Between CLA EB1 BY Y2 Variances EB1	pulation 1.000	Average 0.9974	Std. Dev.	Average	M. S. E.	Cover	
E BY Y2 Variances E Residual Varian Y1 • Between CLU EB1 BY Y2 Variances	1.000	0.9974	0.0200				
Y2 Variances E Residual Varian Y1 • Between CLU EB1 BY Y2 Variances		0.9974	0.0200				
Variances E Residual Varian Y1 • Between CLU EB1 BY Y2 Variances		0.9974	0.0200				
E Residual Varian Y1 • Between CLU EB1 BY Y2 Variances	0.800			0.0203	0.0004	0.950	1.000
Residual Varian Y1 • Between CLU EB1 BY Y2 Variances	0.800						
Y1 • Between CLU EB1 BY Y2 Variances	10000000000	0.7999	0.0232	0.0246	0.0005	0.960	1.000
• Between CLU EB1 BY Y2 Variances	ces						
EB1 BY Y2 Variances	1.200	1.2031	0.0201	0.0204	0.0004	0.930	1.000
Y2 Variances	STER1 Lev	el					
Variances							
	1.000	1.0014	0.1620	0.1720	0.0260	0.940	1.000
EB1							
	0.200	0.2070	0.0477	0.0488	0.0023	0.970	1.000
Residual Varian	ces						
Yl	0.900	0.8982	0.0528	0.0542	0.0028	0.940	1.000
• Between CLU	STER2 Lev	el					
EB2 BY							
¥2	1.000	1.0205	0.1494	0.1413	0.0225	0.960	1.000
Intercepts							
¥1	1.000	0.9868	0.1082	0.1098	0.0118	0.960	1.000
Variances							
EB2	0.600	0.5943	0.1556	0.1450	0.0240	0.920	1.000
Residual Varian							

Tihomir Asparouhov and Bengt Muthén

Muthén & Muthén 16/57

# 3-level SEM Example 2: Factor Analysis Results, 40 level 3 clusters

MODEL RESULTS

			ESTIMATES		S. E.	M. S. E.	95%	% Sig
21010		Population	Average	Std. Dev.	Average		Cover	Coeff
Within	Level							
E	BY							
¥2		1.000	0.9983	0.0273	0.0281	0.0007	0.950	1.000
Varian	ces							
E		0.800	0.7966	0.0312	0.0347	0.0010	0.970	1.000
Residu	al Var	ciances						
Y1		1.200	1.2012	0.0273	0.0284	0.0007	0.950	1.000
Between	CLUST	TER1 Level						
EB1	BY							
¥2		1.000	1.0305	0.2954	0.2631	0.0873	0.930	1.000
Varian	ces							
EB1		0.200	0.2146	0.0819	0.0693	0.0069	0.900	0.910
Residu	al Var	ciances						
¥1		0.900	0.8847	0.0742	0.0765	0.0057	0.960	1.000
Between	CLUST	TER2 Level						
EB2	BY							
¥2		1.000	1.0198	0.1996	0.2088	0.0398	0.950	1.000
Interc	epts							
¥1		1.000	0.9838	0.1579	0.1535	0.0249	0.950	1.000
Varian	ces							
EB2		0.600	0.5848	0.2036	0.2001	0.0413	0.920	0.950
Residu	al Var	ciances						
¥1		0.300	0.2819	0.1076	0.0943	0.0118	0.810	0.900

# 3-level SEM Example 2: Factor Analysis Results, 20 level 3 clusters

MODEL RESULTS

Π

	Population	ESTIMATES Average	Std. Dev.	S. E. Average	M. S. E.	95% Cover	
Within Leve	1	,		-			
E B	Y						
¥2	1.000	0.9948	0.0388	0.0387	0.0015	0.950	1.000
Variances							
E	0.800	0.7974	0.0455	0.0482	0.0021	0.920	1.000
Residual V							
Y1	1.200	1.1977	0.0382	0.0388	0.0015	0.970	1.000
Between CLU	STER1 Level						
EB1 B							
¥2	1.000	1.1296	0.5277	0.4536	0.2925	0.940	0.830
Variances							
EB1	0.200	0.2172	0.1165	0.1002	0.0137	0.920	0.530
Residual V							
¥1	0.900	0.8722	0.1221	0.1104	0.0155	0.930	1.000
Between CLU	STER2 Level						
EB2 B	Y						
¥2	1.000	1.0622	0.3626	0.3714	0.1340	0.900	0.890
Intercepts							
¥1	1.000	0.9687	0.2032	0.2147	0.0418	0.950	0.990
Variances							
EB2	0.600	0.5824	0.2963	0.2640	0.0872	0.830	0.630
Residual V	ariances						
¥1	0.300	0.2742	0.1450	0.1292	0.0215	0.880	0.610

# 3-level SEM Example 2: Factor Analysis Results with random coefficients

• Simulation: 5 observed indicator variables *Y*<sub>pijk</sub> and one covariate *X*<sub>ijk</sub>.

$$egin{aligned} Y_{pijk} &= \mu_p + \lambda_p \eta_{ijk} + arepsilon_{pijk} + arepsilon_{pijk} &= \eta_{0ijk} + s_{jk} x_{ijk} \ \eta_{0ijk} &= \eta_{1ijk} + \eta_{2jk} + \eta_{3k} \ s_{jk} &= s_{2jk} + s_{3k} \end{aligned}$$

- 5 indicators variables measuring one factor regressed on a single predictor with random intercept and slope varying over level 2 and level 3
- 100 level 3 clusters, each with 20 level 2 clusters, each with 5 observations with 5 indicator variables

### 3-level SEM Example 2: Factor Analysis Results with random coefficients results

MODEL RESULTS

			ESTIMATES	3	S. E.	M. S. E.	95%	& Sig
		Population	Average	Std. Dev.	Average		Cover	Coeff
Within	Level							
F	BY							
¥2		1.000	0.9998	0.0165	0.0143	0.0003	0.900	1.000
Residu	al Vari	iances						
¥1		2.000	2.0027	0.0516	0.0480	0.0026	0.930	1.000
F		1.000	0.9989	0.0414	0.0396	0.0017	0.950	1.000
Betweer	CLUSTI	ER1 Level						
S	WITH	н						
F1		0.300	0.2972	0.0329	0.0359	0.0011	0.960	1.000
Variar	ices							
F1		0.700	0.6971	0.0516	0.0492	0.0026	0.930	1.000
S		0.600	0.5926	0.0460	0.0517	0.0022	0.960	1.000
Between	LUSTI	ER2 Level						
S	WITH	H						
F2		0.300	0.2934	0.0601	0.0630	0.0036	0.950	1.000
Means								
s		1.300	1.2985	0.0705	0.0720	0.0049	0.940	1.000
Interd	cepts							
¥1		0.000	-0.0015	0.0769	0.0787	0.0059	0.950	0.050
Variar	ices							
F2		0.500	0.4871	0.0784	0.0822	0.0062	0.950	1.000
S		0.400	0.3981	0.0691	0.0717	0.0047	0.960	1.000

Tihomir Asparouhov and Bengt Muthén Muthén & Muthén 20/57

Three-level structural equation modeling for categorical and continuous variable: Bayesian estimation

### **Bayesian Estimation**

- The same 3-level structural model as with ML
- MCMC estimation: slopes / variances / latent variables / missing data / between level data
- The key is the likelihood posteriors  $[Y_{3k}|*]$  and  $[Y_{2jk}|Y_{3k},*]$  the same as in the EM algorithm, followed by multiple group analysis
- Algorithm uses largest blocks possible for efficient mixing
- Advantages over ML: accommodates priors, not reliant on asymptotic theory, posterior for random effects and more accurate estimates, flexible modeling in structural part
- Easily extends modeling to categorical variables
- New feature in Mplus 7

### 3-level with categorical data: univariate regression

- $y_{ijk}$  categorical variable with 3 possible outcomes
- $x_{1ijk}, x_{2jk}, x_{3k}$  covariates at level 1, 2 and 3.

$$y_{ijk} = l \Leftrightarrow \tau_l < y_{ijk}^* < \tau_{l+1}$$

$$y_{ijk}^* = \alpha_{jk} + \alpha_k + \beta_1 x_{1ijk} + \beta_2 x_{2jk} + \beta_3 x_{3k} + \varepsilon_{ijk}$$

$$\beta_1 = \beta_{10} + \beta_{1jk} + \beta_{1k}$$

$$\beta_2 = \beta_{20} + \beta_{2k}$$

- Variance of  $\varepsilon_{ijk}$  is 1 as in probit regression
- ML would use 5 dimensional integration very difficult
- Bayesian estimation  $\leq 1 \text{ min}$
- Default of weakly informative priors: IW(I,p+1), N(0,5). Key for small sample problems.

### 3-level with categorical data: univariate regression results

Between CLUSTER1 Level

S	ON							
x2	011	0.400	0.3938	0.0635	0.0550	0.0040	0.880	1.000
Y	WITH							
s	with	1.000	0.9732	0.1201	0.1151	0.0150	0.880	1.000
Residu	al Varian	ces						
Y		1.500	1.4758	0.1676	0.1478	0.0284	0.910	1.000
S		1.600	1.5260	0.1850	0.1740	0.0393	0.920	1.000
Between	CLUSTER2	Level						
S	ON							
X3		0.550	0.5183	0.2106	0.2246	0.0449	0.950	0.620
82	ON							
x3		0.550	0.5646	0.2540	0.2135	0.0641	0.870	0.710
Y	ON							
X3		1.300	1.2793	0.1815	0.1795	0.0330	0.920	1.000
S2	WITH							
S		0.300	0.2596	0.2505	0.2412	0.0637	0.930	0.220
Y		-0.200	-0.2188	0.1844	0.1908	0.0340	0.920	0.180
Y	WITH							
S		0.700	0.6413	0.2068	0.2182	0.0458	0.950	0.940
Interc	cepts							
S		1.700	1.6872	0.1582	0.1573	0.0249	0.940	1.000
52		0.300	0.2971	0.1497	0.1487	0.0222	0.930	0.540
Thresh	nolds							
Y\$1		-2.100	-2.0968	0.1380	0.1152	0.0189	0.860	1.000
¥\$2		2.100	2.0661	0.1600	0.1179	0.0265	0.820	1.000
Residu	al Varian	ces						
Y		1.300	1.2368	0.2444	0.2315	0.0631		
S		2.200	2.0684	0.3702	0.3756	0.1530		
S2		2.000	1.9354	0.3586	0.3428	0.1314	0.950	1.000
th o main	A		nat Muthán	3.5.0	in & Muth	/	157	

Tihomir Asparouhov and Bengt Muthén Muthén & Muthén

24/57

### 3-level SEM with categorical data: Factor model

- $y_{pijk}$  5 categorical variable with 3 possible outcomes
- 1 factor on each level
- ML would use 13 dimensions of integration

$$y_{pijk} = l \Leftrightarrow \tau_{pl} < y_{pijk}^* < \tau_{pl+1}$$

$$y_{pijk}^* = y_{pjk} + y_{pk} + \lambda_{1p} f_{1ijk} + \lambda_{2p} f_{2jk} + \lambda_{3p} f_{3k} + \varepsilon_{pijk}$$

• Variances of  $\varepsilon_{pijk}, f_{1ijk}, f_{2jk}, f_{3k}$  are 1

### 3-level SEM with categorical data: Factor model results

MODEL RESULTS

Within Leve	Population	ESTIMATES Average	Std. Dev.	S. E. Average	M. S. E.	95% Cover	% Sig Coeff
E1 H Y1	1.300	1.3130	0.0566	0.0496	0.0033	0.890	1.000
Between CLU	JSTER1 Level						
E2 I Y1	1.000	1.0094	0.0639	0.0564	0.0041	0.920	1.000
Residual V Y1	Variances 0.400	0.4042	0.0607	0.0573	0.0037	0.930	1.000
Between CLU	JSTER2 Level						
E3 F Y1	3Y 0.600	0.6032	0.1056	0.1072	0.0111	0.940	1.000
Thresholds Y1\$1 Y1\$2	-1.000	-0.9994	0.1077	0.0866	0.0115		
Residual V Y1	Variances 0.300	0.3139	0.0750	0.0772	0.0058	0.960	1.000

# Cross-classified structural equation modeling

### Cross-classified data

- *Y<sub>pijk</sub>* is the *p*-th observation for person *i* belonging to level 2 cluster *j* and level 3 cluster *k*.
- Level 2 clusters are not nested within level 3 clusters
- Examples:
  - To model income: individuals are nested within the same geographical location and are nested within occupation clusters
  - Students are nested within schools and nested within neighborhoods
  - Student performance scores are nested within students and within teachers
  - Studies where observations are nested within persons and treatments/situations
  - Studies where observations are nested within neighborhoods and interviewer
  - Generalizability theory (Brennan, 2001; Cronbach, Rajaratnam, & Gleser, 1963), Items are considered a random sample from a population of items.

### Cross-classified modeling

- Why do we need to model both sets of clustering?
- A model that ignores the dependence between observations in the same cluster yields misspecifications, underestimates SE, fails to discover the true predictor/explanotory effect stemming from the clusters
- Why do we need random effects and not fixed effects?
- Fixed effects possible: Tucker3
- Fixed effect for one set of clusters and random for the other: dummy variables for one set of clusters
- If the number of clusters is more than 10 modeling with fixed effects may lead to too many parameters and less accurate model.

#### Gonzalez, De Boeck, Tuerlinckx (2008) A Double-Structure Structural Equation Model for Three-Mode Data. Psychological Methods, 337 - 353.

Tucker3 (a) Variable Tucker3 (b) SEM-MTMM Interactions [PSR] [PSR] [PS][PR] [PR1[SR] Parameters Persons fixed random random random Situations fixed fixed fixed random Responses fixed fixed Dependence structure Dependent irrelevant pairs (S, R) pairs (S, R) R Independent irrelevant nonoverlapping pairs (P, S)

Table 1 Three Approaches for Modeling Three-Mode Data

Note: Variables that interact are grouped together inside a set of brackets; variables that are independent of one another are placed in separate brackets. SEM\_MTMM = structural equation model for multitrait-multimethod data; 2sSEM = double-structure structural equation model; P = persons; S = situations; R = responses.

- Univariate model
- Both cluster levels contribute with random effects

$$Y_{ijk} = Y_{1ijk} + Y_{2j} + Y_{3k}$$

•  $Y_{2j}$  and  $Y_{3k}$  are random effects for the two cluster levels

### Cross-classified model

• General SEM model

$$Y_{pijk} = Y_{1pijk} + Y_{2pj} + Y_{3pk}$$

• 3 sets of structural equations - one on each level

$$egin{aligned} Y_{1ijk} &= \mathbf{v} + \Lambda_1 \eta_{ijk} + arepsilon_{ijk} \ &= lpha + B_1 \eta_{ijk} + \Gamma_1 x_{ijk} + \xi_{ijk} \ &Y_{2j} &= \Lambda_2 \eta_j + arepsilon_j \ &\eta_j &= B_2 \eta_j + \Gamma_2 x_j + \xi_j \end{aligned}$$

$$Y_{3k} = \Lambda_3 \eta_k + \varepsilon_k$$
$$\eta_k = B_3 \eta_k + \Gamma_3 x_k + \xi_k$$

- The parameter v and  $\alpha$  can be used on any level
- Bayesian MCMC estimation
- The two sets of random effects  $Y_{2pj}$ ,  $Y_{3pk}$  are treated separately
- The two sets of random effects  $[Y_{2pj}|*, Y_{3pk}]$  and  $[Y_{3pk}|*, Y_{2pj}]$  are level 2 random effect
- Easily extends to categorical variables
- New feature in Mplus 7

### Cross-classified model, example 1: Factor model

• 1 factor at the individual level and 1 factor at each of the clustering levels, 5 indicator variables on the individual level

$$y_{pijk} = \mu_p + y_{pj} + y_{pk} + \lambda_{1p}f_{1ijk} + \lambda_{2p}f_{2j} + \lambda_{3p}f_{3k} + \varepsilon_{pijk}$$

- Variances of  $f_{1ijk}, f_{2jk}, f_{3k}$  are 1, all loadings are estimated.
- 50 level 2 clusters, 50 level 3 clusters, 1 unit within each cluster intersection
- Total sample size 2500
- 1 unit on the within level would not work if the clusters were nested
- Estimation takes less than 1 min per replication

### Cross-classified model example 1: Factor model results

MODEL RESULTS

Within	Level	Population	ESTIMATES Average	Std. Dev.	S. E. Average	M. S. E.		% Sig Coeff
E	BY							
¥1		1.500	1.5018	0.0286	0.0324	0.0008	0.970	1.000
Residu	al Var	ciances						
¥1		1.200	1.1999	0.0400	0.0437	0.0016	0.990	1.000
Betweer	LUST	TER1 Level						
E2	BY							
¥1		1.000	1.0000	0.2822	0.2865	0.0788	0.920	0.940
Residu	al Var	ciances						
¥1		1.500	1.6146	0.3755	0.4754	0.1527	0.990	1.000
Betweer	1 CLUSI	TER2 Level						
E3	BY							
¥1		0.800	0.8606	0.1538	0.1698	0.0271	0.950	1.000
Interd	cepts							
¥1	•	2.200	2.1561	0.2477	0.2900	0.0627	0.990	1.000
Residu	al Var	ciances						
¥1		0.500	0.5517	0.1582	0.1647	0.0274	0.940	1.000
		0.000	0.0017	0.1002	0.1017	0.0271	0.0.0	1.000

# Cross-classified model, example 2: Simulation on Gonzalez's example

- Observations are nested within individual and several specific situations, the dependent variables are 4 binary outcomes
- 1 observation for each pair of clustering units

$$y_{pjk}^* = y_{pj} + y_{pk} + \varepsilon_{pjk}$$

- Variances of  $\varepsilon_{pjk}$  is fixed to 1
- 100 level 2 clusters, 100 level 3 clusters
- Identical structural model for the two cluster random effects

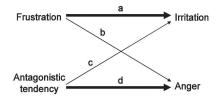
$$y_{1j} = \beta_1 y_{3j} + \beta_2 y_{4j} + \varepsilon_{1j}$$
  

$$y_{2j} = \beta_3 y_{3j} + \beta_4 y_{4j} + \varepsilon_{2j}$$
  

$$y_{1k} = \beta_1 y_{3k} + \beta_2 y_{4k} + \varepsilon_{1k}$$
  

$$y_{2k} = \beta_3 y_{3k} + \beta_4 y_{4k} + \varepsilon_{2k}$$

## Cross-classified model, example 2: Simulation on Gonzalez's example



*Figure 3.* Graphical representation of the research questions. a, b, c, and d are effect parameters.

Identical structural model is estimated for the two-sets of random effects.

## Cross-classified model, example 2: Simulation on Gonzalez's example results

Between CLUSTER1 Level

¥1	ON							
Z1		0.300	0.3059	0.0891	0.0932	0.0079	0.940	0.910
22		-0.300	-0.2932	0.0870	0.0900	0.0075	0.970	0.940
¥2	ON							
Z1		0.500	0.5046	0.0834	0.0877	0.0069	0.970	1.000
22		-0.500	-0.5045	0.1004	0.0939	0.0100	0.960	1.000
¥1	WITH							
Y2		0.700	0.6823	0.1553	0.1443	0.0242	0.940	1.000
21	WITH							
Z2		0.700	0.7129	0.1394	0.1458	0.0194	0.970	1.000
Varian	ices							
Z1		1.500	1.5237	0.2282	0.2370	0.0521	0.960	1.000
Z2		0.900	0.9117	0.1251	0.1408	0.0156	0.960	1.000
Residu	al Varian	ices						
¥1		1.500	1.4707	0.2335	0.2317	0.0548	0.950	1.000
¥2		0.900	0.8946	0.1421	0.1427	0.0200	0.970	1.000

## Cross-classified model, example 2: Simulation on Gonzalez's example results

Between CLUSTER2 Level

¥1	ON							
21		0.300	0.3059	0.0891	0.0932	0.0079	0.940	0.910
22		-0.300	-0.2932	0.0870	0.0900	0.0075	0.970	0.940
¥2	ON							
21		0.500	0.5046	0.0834	0.0877	0.0069	0.970	1.000
Z2		-0.500	-0.5045	0.1004	0.0939	0.0100	0.960	1.000
¥1	WITH							
¥2		0.400	0.3929	0.0788	0.0814	0.0062	0.960	1.000
21	WITH							
22		0.400	0.4056	0.0805	0.0816	0.0064	0.950	1.000
Thresh	olds							
¥1\$1		0.200	0.1998	0.1400	0.1461	0.0194	0.970	0.310
¥2\$1		-0.500	-0.4932	0.1400	0.1453	0.0194	0.950	0.910
Z1\$1		0.200	0.2263	0.1517	0.1398	0.0235	0.930	0.350
Z2\$1		-0.500	-0.4675	0.1503	0.1285	0.0234	0.880	0.910
Variar	ices							
21		0.500	0.5208	0.0767	0.0811	0.0063	0.950	1.000
22		0.800	0.8133	0.1342	0.1255	0.0180	0.960	1.000
Residu	al Varian	ices						
¥1		0.500	0.5053	0.0748	0.0802	0.0056	0.950	1.000
¥2		0.800	0.7995	0.1277	0.1265	0.0161	0.960	1.000

# Multiple imputations for two-level data

#### Multiple imputations for two-level data

• Bayesian estimation of unrestricted H1 model

$$y_{pij} = y_{wpij} + y_{bpj}$$

- Unrestricted variance covariance matrix for y<sub>wpij</sub>
- Unrestricted means and variance covariance matrix for y<sub>bpj</sub>
- Imputation of data based on the MCMC estimation
- Categorical data imputation based on *y*<sup>\*</sup><sub>*pij*</sub> and unrestricted correlation matrices
- PX methodology for estimating correlation matrices: Mplus 6.1

#### Multiple imputations for two-level categorical data example

- Asparouhov and Muthén (2010) Multiple Imputation with Mplus
- P categorical variables with 4 categories, M clusters of size 30
- Two-level factor model with 1 factor at each level

$$Y_{pij}^{*} = \lambda_{wp} \eta_{wij} + \lambda_{bp} \eta_{bj} + \varepsilon_{bpj} + \varepsilon_{wpij}$$

- Variance of  $\varepsilon_{wpij}$  is fixed to 1
- Generate MAR missing data using

$$P(Y_j \text{ is missing}) = \frac{1}{1 + Exp(-1.5 + 0.15\sum_{k=26}^{30} Y_k)}.$$

• Analyze the data with WLSMV directly and WLSMV after imputation

Table: *MSE* for the threshold parameters for two-level imputation with categorical variables.

Number of	Number of	Direct	H1 Imputed
Clusters	Variables	WLSMV	WLSMV
50	10	0.352	0.268
200	10	0.204	0.077
50	30	0.418	0.273
200	30	0.235	0.074

Imputation based results are more accurate. WLSMV supports MCAR, MARX but not MAR.

- Mplus 5
- Based on ML estimation for all continuous variables and WLS for continuous and categorical
- Estimate an unrotated solution

$$Y_{ij} = \mu + \Lambda_{0w} \eta_{0w} + \Lambda_{0b} \eta_{0b} + \varepsilon_{ij}$$

- $\Lambda_{0w}$  and  $\Lambda_{0b}$  have fixed 0 above the diagonal
- $\eta_{0w}$  and  $\eta_{0b}$  are standard normal
- Independent rotation of  $\Lambda_{0w}$  and  $\Lambda_{0b}$  based on Jennrich's gradient projection algorithm (GPA)
- GPA minimizes  $f(\Lambda_{0w}H_w^{-1})$  and  $f(\Lambda_{0b}H_b^{-1})$  over orthogonal rotations  $H_w$  and  $H_b$  where f is the rotation criteria

• Mplus uses as a default the Geomin rotation criteria

$$f(\Lambda) = \sum_{i=1}^{p} \left( \prod_{j=1}^{m} (\lambda_{ij}^2 + \varepsilon) \right)^{1/m}$$

- Chi-square test of fit for evaluating the model
- As a preliminary step in deciding the number of factors estimate within level EFA or between level EFA
- Within level EFA estimates unrestricted variance covariance on the between level and factor model on the within level only
- Between level EFA estimates estimates unrestricted variance covariance on the within level and a factor model on the between level only

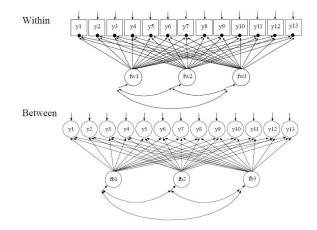
#### **Exploratory Factor Analysis Of Aggression Items**

	Almost Never (scored as 1)	Rarely (scored as 2)	Sometimes (scored as 3)	Often (scored as 4)	Very Often (scored as 5)	Almost Always (scored as 6)
Stubborn	42.5	21.3	18.5	7.2	6.4	4.1
Breaks Rules	37.6	16.0	22.7	7.5	8.3	8.0
Harms Others	69.3	12.4	9.40	3.9	2.5	2.5
Breaks Things	79.8	6.60	5.20	3.9	3.6	0.8
Yells at Others	61.9	14.1	11.9	5.8	4.1	2.2
Takes Others' Property	72.9	9.70	10.8	2.5	2.2	1.9
Fights	60.5	13.8	13.5	5.5	3.0	3.6
Harms Property	74.9	9.90	9.10	2.8	2.8	0.6
Lies	72.4	12.4	8.00	2.8	3.3	1.1
Talks Back to Adults	79.6	9.70	7.80	1.4	0.8	1.4
Teases Classmates	55.0	14.4	17.7	7.2	4.4	1.4
Fights With Classmates	67.4	12.4	10.2	5.0	3.3	1.7
Loses Temper	61.6	15.5	13.8	4.7	3.0	1.4

Item Distributions for Cohort 3: Fall 1st Grade (n=362 males in 27 classrooms)

#### Hypothesized Aggressiveness Factors

- Verbal aggression
  - Yells at others
  - Talks back to adults
  - Loses temper
  - Stubborn
- · Property aggression
  - Breaks things
  - Harms property
  - Takes others' property
  - Harms others
- Person aggression
  - Fights
  - Fights with classmates
  - Teases classmates



Number of clusters 27

Average cluster size 13.407

#### Estimated Intraclass Correlations for the Y Variables

	Intraclass		Intraclass		Intraclass
Variable	Correlation	Variable	Correlation	Variable	Correlation
U1	0.110	U2	0.121	U3	0.208
U4	0.378	U5	0.213	U6	0.250
U7	0.161	U8	0.315	U9	0.208
U10	0.140	U11	0.178	U12	0.162
U13	0.172				

Within-level	Between-level				
Factors	Factors	Df	Chi-Square	CFI	RMSEA
unrestricted	1	65	66(p=0.43)	1.000	0.007
1	1	130	670	0.991	0.107
2	1	118	430	0.995	0.084
3	1	107	258	0.997	0.062
4*	1	97	193	0.998	0.052

\*4th factor has no significant loadings

	Within-Level Loadings			Between-Level Loadings	
	Property	Verbal	Person	General	
Stubborn	0.00	0.78*	0.01	0.65*	
Breaks Rules	0.31*	0.25*	0.32*	0.61*	
Harms Others and Property	0.64*	0.12	0.25*	0.68*	
Breaks Things	0.98*	0.08	-0.12*	0.98*	
Yells At Others	0.11	0.67*	0.10	0.93*	
Takes Others' Property	0.73*	-0.15*	0.31*	0.80*	
Fights	0.10	0.03	0.86*	0.79*	
Harms Property	0.81*	0.12	0.05	0.86*	
Lies	0.60*	0.25*	0.10	0.86*	
Talks Back To Adults	0.09	0.78*	0.05	0.81*	
Teases Classmates	0.12	0.16*	0.59*	0.83*	
Fights With Classmates	-0.02	0.13	0.88*	0.84*	
Loses Temper	-0.02	0.85*	0.05	0.87*	

### Two-level multiple group modeling

#### Two-level multiple group modeling

- Two types of grouping variable: between (private v.s. public schools), within (female v.s. males)
- Between grouping variable: straight forward, likelihood is sum of independent groups
- Within grouping variable: if you treat it as if the grouping variable is between the two groups within the cluster are independent misspecification
- Modeling options: the between random effect is the same across groups OR more flexible the between random effects is different across groups
- Is teacher effect of student performance the same for females and males?
- Are there two correlated random effects (teachers ability to relate to different gender) or is it one random effect (teachers ability) that may affect students performance differently?

#### Two-level multiple group modeling

• Within level grouping variable - one random effect (assuming fixed variance of 1): twolevel mixture(with known class variable)

$$Y_{ijg} = \mu_g + \beta_g Y_j + \varepsilon_{ijg}$$

• Within level grouping variable - two random effects (assuming different variance across groups): twolevel mixture(with known class variable)

$$Y_{ijg} = \mu_g + Y_{gj} + \varepsilon_{ijg}$$

- Both use numerical integration currently
- Using 3 level modeling and treating the grouping variable as level 2 variable ( $Y_{gj}$  are correlated via  $Y_j$ ) : no numerical integration

$$Y_{ijg} = \mu_g + Y_{gj} + Y_j + \varepsilon_{ijg}$$

• LRT to decide between the three models - if nested

### **Plausible Values**

#### **Plausible Values**

- With Baysian estimation we can obtain posterior distribution for the random effects and plausible valeus as draws from the posterior
- Plausible values are more accurate than ML factor scores especially for small sample size
- Plausible values variability accounts for the variability in the parameter estimates
- Can be used for further analysis comparing the random effects of different clusters, estimate difference, can be used for further modeling
- In 3 level modeling plausible values on level 2 accurately reproduce the correlation between level 2 random effects, not factor scores.