## Using Mplus To Investigate Direct Effects in Latent Class Analysis

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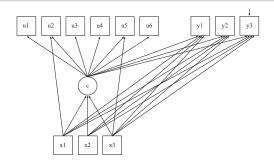
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We thank Thuy Nguyen and Noah Hastings for expert assistance.

#### Outline

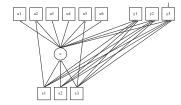
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#### Direct Effects from Covariates to Latent Class Indicators



- Two direct effects: From x1 to u2 and from x3 to u5
- A direct effect implies that for a given class, the probability of u is not the same for different values of x
  - The measurement model parameters are not the same for different individuals - referred to as measurement non-invariance, differential item functioning, item bias
  - For example, with a binary x describing two groups, the measurement instrument does not work the same for the groups

### Direct Effects: Analysis Impact



- Consequences of ignoring direct effects:
  - Violation of conditional independence given C
    - Class enumeration for measurement model impact likely small
  - Distorted model estimates:
    - Class probabilities typically small impact (with or without X)
    - C ON X large impact (direct effects are forced to go through only C (Asparouhov & Muthén, 2014; Web Note 15)
    - Y ON X some impact (when C ON X changes, the indirect effects of X on Y via C change and therefore Y ON X)
- Multistep analysis
  - Let the measurement model include the X part and its direct effects on the latent class indicators, then add distal outcomes

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### Direct Effects: Antisocial Behavior Example

#### • ASB data:

- 17 antisocial behavior items collected in the 1980 National Longitudinal Survey of Youth for respondents between the ages of 16 and 23 together with a set of background variables
- The ASB items assessed the frequency of various behaviors during the past year, here dichotomized as 0 vs > 0 times
- A sample of 7,326 respondents has complete data on the antisocial behavior items and the background variables

#### • ASB analyses:

- SEM (MIMIC) 4-factor analysis (Muthén, 2025)
- Latent class analysis from Mplus Short Course Topic 5, slides 91-118
  - 4-class and 5-class LCA of the 17 latent class indicators
  - 5 classes: High, property offense, drug, person offense, normative (low, except for pot)
  - 4 classes used in this talk
- ASB is a general population survey so that considerable heterogeneity among individuals can be expected - direct effects
- 17U, 11X version and 7U, 7X reduced version with 2 distals

## Input for ASB Analysis with 17 U's and 11 X's C ON X but No Direct Effects

TITLE: ANALYSIS: TYPE = MIXTURE;

DATA: FILE = asbfree.dat; STARTS = 400 100; FORMAT = 34X 54F2.0; PROCESSORS = 12;

VARIABLE: NAMES = property fight shoplift lt50 MODEL: %OVERALL%

gt50 force threat injure pot drug c ON sex-abuse; soldpot solddrug con auto bldg goods

gambling dsm1-dsm22 sex black hisp single divorce dropout college onset f1 f2 f3 age94 cohort dep abuse:

USEVARIABLES = property-gambling sex black hisp single divorce dropout college onset age94 dep abuse;

CATEGORICAL = property-gambling; CLASSES = c(4);

• 187 possible direct effects - which ones are important to include?

ESTIMATOR = ML:

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### Approaches for Searching for Direct Effects

- Analysis with latent class variable regressed on all covariates and in addition:
  - Each latent class indicator regressed on all covariates or All latent class indicators regressed on one covariate
    - Pro: Each analysis easily converges
    - Con: Several analyses one for each latent class indicator/covariate
  - All latent class indicators regressed on all covariates
    - Pro: Single analysis
    - Con: May not converge or be empirically identified
       Relies on higher-order moments (cf. the non-identified case of MIMIC with direct effects for continuous factor indicators)
  - All latent class indicators regressed on all covariates using PSEM regularization, Asparouhov & Muthén (2024).
    - PSEM for mixtures: Asparouhov & Muthén (2025)
      - Pro: Leads to parsimonious models, i.e., fewer significant direct effects, converges more easily for a small-enough prior variance
      - Con: Prior variance choice calls for more than one analysis

### PSEM Applied to Finding Direct Effects in LCA

- PSEM: Penalized structural equation modeling
  - ML estimation using priors
- PSEM uses priors for two purposes:
  - Estimating models that are non-identified without the priors
    - Similar to BSEM for Bayes estimation
  - Simplifying models that are identified but can be fitted practically as well with fewer parameters
    - Reguralized analysis (RegSEM). Common prior: LASSO
    - Example: Direct effects in LCA
- Including all direct effects, PSEM uses the LASSO or ALF (Alignment Loss Function) mean and variance priors for the directs effects, maximizing: fit function = log likelihood + penalty
  - where the penalty is larger for smaller variance. Ex: ALF(0, 0.01)
    - The penalty (which is negative) penalizes models with many direct effects - favors a parsimonious model
      - Variance = 0: Same as non-PSEM analysis with no direct effects
      - Variance = ∞: Same as non-PSEM analysis with all direct effects
    - Goal: use a variance that makes the logL practically as good as with all direct effects included - but with fewer direct effects

### Input for PSEM with ALF(0, 0.5): 17 U's, 11 X's

TITLE: 17 U's, 11 X's,

PSEM ALF(0.5)

Model 6 on slide 12

DATA: FILE = asbfree.dat:

VARIABLE: NAMES = property fight shoplift 1t50

gt50 force threat injure pot drug soldpot solddrug con auto bldg goods gambling dsm1-dsm22 sex black hisp single divorce dropout college onset f1 f2 f3

age94 cohort dep abuse;

USEVAR = property-gambling sex black hisp single divorce dropout

CATEGORICAL = property-gambling;

CLASSES = c(4);

college onset age94 dep abuse;

ANALYSIS: TYPE = MIXTURE:

ESTIMATOR - ML:  $STARTS = 800\ 200$ 

PROCESSORS = 12:

MODEL: %OVERALL%

> c ON sex-abuse: ! Direct effects:

property-gambling ON

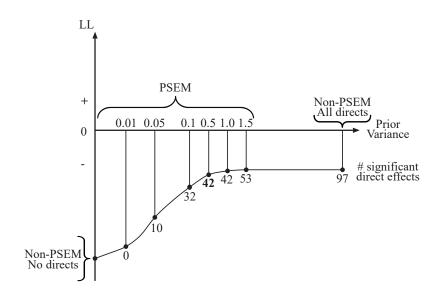
sex-abuse (d1-d187);

MODEL PRIOR:  $d1-d187\sim ALF(0,0.5);$ 

OUTPUT: TECH1 TECH3 TECH10

SVALUES:

### ASB Log Likelihoods: 17 Us, 11 Xs (187 Possible Effects)



#### ASB Log Likelihood and BIC Values for 17 Us, 11 Xs

Model	# par's	logL	BIC	#sig. dir.	logL drop
1. No directs	104	-40,088.255	81,102	0	
2. All directs	291	-39,278.105	81,146	97	
3. 97 sig. directs	201	-39,388.584	80,466	97	
4. PSEM (0.05)	NA	-39,564.585	NA	10	
5. PSEM (0.1)	NA	-39,384.232	NA	32	
6. <b>PSEM (0.5)</b>	NA	-39,295.792	NA	42	
7. <b>PSEM (1.0)</b>	NA	-39,282.084	NA	42	
8. PSEM (1.5)	NA	-39,279.840	NA	53	
9. 4: 10 directs	114	-39,855.419	80,725	10	1.2 %
10. 5: 32 directs	136	-39,531.593	80,273	32	0.4 %
11. 6: 42 directs	146	-39,466.120	80,232	41	0.2 %
12. 7: 42* directs	146	-39,485.213	80,270	41	0.2 %
13. 8: 53 directs	157	-39,483.475	80,364	52	0.2 %

logL drop is computed as the percentage 100 \* (logL - logL<sub>Model 3</sub>)/logL<sub>Model 3</sub>

<sup>\*</sup> The 42 direct effects are not all the same in models 11 and 12

# Discussion of ASB Models for 17 Us, 11 Xs: Why is 42 for PSEM(0.5) bolded in the Graph of Slide 12?

- PSEM(0.5) model 6 has a lower (worse) logL than PSEM(1.0) model 7
- Both models show 42 significant direct effects but not the same ones
- The non-PSEM models 11 and 12 are based on PSEM models 6 and 7
  - Freeing the 42 effects results in a higher logL for model 11 than for model 12 despite model 6 having a lower logL than model 7
- Models 11 and 12 have the same
  - Number of parameters
  - Number of significant direct effects
  - LogL 0.2 % drop
- Model 11 is chosen
  - Better logL and BIC than model 12
  - Best BIC of all the models

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## ASB Inputs for the Reduced Set of 7 U's, 7 X's DATA and VARIABLE Commands for All Runs

TITLE:

DATA: FILE = asbfree.dat;

VARIABLE: NAMES = property fight shoplift lt50 gt50 force threat injure pot drug soldpot solddrug con auto

bldg goods gambling dsm1-dsm22 sex black hisp single divorce dropout college onset f1 f2 f3 age94

cohort dep abuse;

USEVARIABLES = property fight shoplift threat

pot drug goods

sex black hisp single divorce dropout age94;

CATEGORICAL = property-goods;

CLASSES = c(4);

### All Latent Class Indicators Regressed on All Covariates

#### All U's on all X's, PSEM (1.0)

All U's on all X's

ANALYSIS: TYPE = MIXTURE;

ANALYSIS: TYPE = MIXTURE: ESTIMATOR = ML:

YSIS: TYPE = MIXTURE; ESTIMATOR = ML; STARTS = 400 100;

 $STARTS = 400\ 100;$  PROCESSORS = 8;

PROCESSORS = 8;

MODEL: %OVERALL% C ON sex-age94

%OVERALL% c ON sex-age94; c ON sex-age94; property-goods ON

property-goods ON sex-age94; sex-age94 (d1-d49);

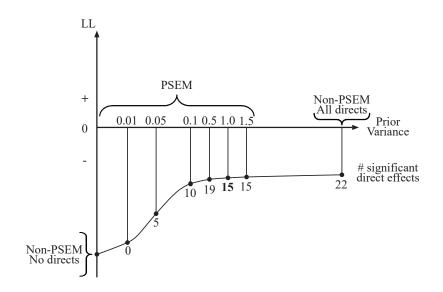
OUTPUT: TECH10 SVALUES; OUTPUT: TECH10 SVALUES;

MODEL PRIOR:  $d1-d49 \sim ALF(0,1.0)$ ;

! Computing time: 1 minute

! (17U-11X run takes 11 minutes)

### ASB Log Likelihoods: 7 Us, 7 Xs (49 Possible Effects)



### ASB Log Likelihood and BIC Values for 7 Us, 7 Xs

Model	# par's	logL	BIC	#sig. dir.	logL drop
1. No directs	52	-23,930.193	48,323	0	
2. All directs	101	-23,653.011	48,205	22	
3. 22 sig. directs	74	-23,670.477	47,999	22	
4. PSEM (0.1)	NA	-23,680.907	NA	10	
5. PSEM (0.5)	NA	-23,656.804	NA	19	
6. <b>PSEM (1.0)</b>	NA	-23,654.305	NA	15	
7. PSEM (1.5)	NA	-23,653.627	NA	15*	
8. 4: 10 directs	62	-23,774.174	48,100	10	0.4 %
9. 5: 19 directs	71	-23,683.270	47,998	19	0.1 %
10. 6: 15 directs	67	-23,715.072	48,026	14	0.2 %

- logL drop is computed as the percentage  $100 * (logL logL_{Model 3})/logL_{Model 3}$
- \* The 15 direct effects are the same in models 6 and 7

## Discussion of ASB Models for 7 Us, 7 Xs: Why is 15 for PSEM(1.0) bolded in the Graph of Slide 18?

- PSEM(1.0) model 6 and PSEM(1.5) model 7 have the same number of significant direct effects
  - It doesn't matter that model 7 has a higher logL because both models lead to the non-PSEM model 10
- Model 10 is chosen because it has a small 0.2 % drop in logL, is relatively parsimonious, and is close to the best BIC
  - Model 9 is a strong contender with a smaller 0.1 % drop in logL, the best BIC, but 19 instead of 15 direct effects to consider
  - Model 3 is also a contender with almost the same BIC as model
     9, but has 22 instead of 15 direct effects
    - Are the extra direct effects of model 3 and model 9 important? <sup>1</sup>
- Class-specific direct effects can be explored based on model 10 using PSEM DIFF priors

<sup>&</sup>lt;sup>1</sup>For model 9, Fight on Dropout is the only noteworthy direct effect beyond those of model 10 (probability difference 0.14 for non-hispanic males at average Age94)

#### Technical Notes

- # parameters = NA because the RegSEM version of PSEM doesn't have a good way of counting the parameters:
  - Parameters can have tiny prior variances with parameters estimated close to 0, and contributing nothing to fit
    - These are not real parameters and shouldn't be counted, but the precise removal of these parameters is subjective since it will need a definition of how small is really 0
- Alternative log likelihood drop % definitions:
  - PSEM models: (L L(M2)) / (L(M1) L(M2))
    - The logic is that the H1 model here is M2 and baseline is M1
  - Models based on PSEM: (L L(M3)) / (L(M1) L(M3))
    - The logic is that the H1 model here is M3 and baseline is M1
  - % drop relative to the total possible drop: Cut-off  $\leq 5\%$ ?
  - For the full set of variables, model 11 does best same model choice as on slides 13-14
  - For the reduced set of variables, model 9 does best with 19 direct effects (5% drop), whereas slides 18-20 chose model 10 with 15 direct effects (17% drop) are the extra direct effects important?

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## Interpreting the 4 Latent Classes for Model 10 with 15 Direct Effects

- Classes (probability):
  - Fight/Threat (0.321), High (0.130), Drugs (0.172), Low (0.377)
- The output section RESULTS IN PROBABILITY SCALE provides convenient interpretation, especially with binary latent class indicators
  - Probability of being in the high category of the indicator:

Variable	Class 1	Class 2	Class 3	Class 4
PROPERTY	0.18	0.72	0.14	0.03
FIGHT	0.49	0.73	0.12	0.02
SHOPLIFT	0.22	0.79	0.39	0.07
THREAT	0.58	0.80	0.30	0.05
POT	0.32	0.90	0.96	0.18
DRUG	0.03	0.53	0.55	0.01
GOODS	0.08	0.53	0.09	0.01

# Class Probabilities for 7 U's on 7 X's Comparing Three Models

- 4 classes: Fight/Threat, High, Drugs, Low
  - Same interpretation for these 3 different models
- Class probabilities with 15 direct effects: 0.321, 0.130, 0.172, 0.377
- Class probabilities with no direct effects: 0.298, 0.126, 0.189, 0.386
- Class probabilities with latent class indicators only: 0.219, 0.106, 0.209, 0.466

## Comparing C ON X Results for 7 U's on 7 X's: 15 Directs (Model 10) vs No Directs for Class 1

15 direct effects, Model 10			No direct effects				
C#1 ON X				C#1 ON X			
	Estimate	S.E.	Est./S.E.		Estimate	S.E.	Est./S.E.
SEX	1.286	0.102	12.627	SEX	1.590	0.101	15.809
BLACK	1.261	0.128	9.834	BLACK	1.088	0.114	9.572
HISP	-0.256	0.128	-1.993	HISP	0.210	0.121	1.734
SINGLE	0.011	0.107	0.105	SINGLE	0.068	0.109	0.620
DIVORCE	0.272	0.120	2.278	DIVORCE	0.141	0.122	1.160
DROPOUT	0.282	0.132	2.139	DROPOUT	0.427	0.131	3.252
AGE94	-0.204	0.024	-8.634	AGE94	-0.316	0.025	-12.673

- Number of significant effects of covariates on all the latent classes:
  - 15 direct effects model: 17
  - No direct effects model: 13
- Similar discrepancies found in the simulations of Asparouhov & Muthén (2014; Web Note 15)

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### 15 Direct Effects for 7 U's on 7 X's (Model 10)

• Property: Sex, Black, Divorce, Age94

• Fight: Sex, Hisp, Age94

• Shoplift: Black, Hisp

• Threat: Black

• Pot: Age94

Drug: Black, Age94

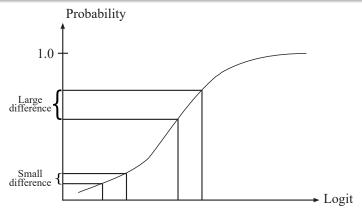
Goods: Sex, Black

## Direct Effect Estimates for Property and Fight Indicators

	Estimate	S.E.	Est./S.E.
PROPERTY ON			
SEX	0.676	0.098	6.925
BLACK	-0.744	0.111	-6.673
DIVORCE	-0.458	0.112	-4.078
AGE94	-0.100	0.021	-4.857
FIGHT ON			
SEX	0.587	0.090	6.501
HISP	0.731	0.125	5.860
AGE94	-0.110	0.019	-5.658

- The sex (male) effect is larger for Property than for Fight
- But the effect on their probabilities also depends on their thresholds

## Direct Effect Probabilities: Prob = $1/(1 + e^{-Logit})$

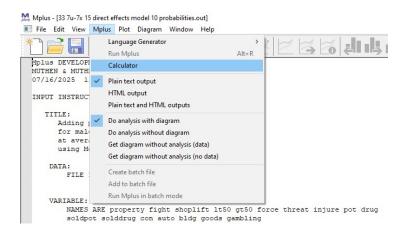


- Logit =  $\tau_c + \beta x$ , where  $\tau_c$  is a threshold for class c and  $\beta$  is the direct effect
- Small probability obtained with large threshold, resulting in small logit
- Largest probability change from x = 0 to x = 1 occurs at the steepest part of the curve with threshold close to zero, that is, probability close to 0.5 (logit = 0 gives probability = 0.5)

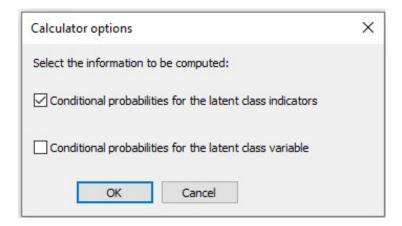
## Direct Effect Probabilities: Property and Fight Indicators Comparing Males and Females for the Fight/Threat Class

- Model 10 estimates:
  - Property:
    - Threshold = 1.352,  $\beta_{male}$  = 0.676 (Logit = -0.68 when other x's=0)
  - Fight:
    - Threshold = 0.217,  $\beta_{male}$  = 0.587 (Logit = 0.37 when other x's=0)
- Probabilities for females (sex=0) vs males (sex=1) at zero values for all other covariates, except age94 which is at its sample mean:
  - Property: 0.161 vs 0.274 (difference = 0.113)
  - Fight: 0.368 vs 0.511 (difference = 0.143)
- Compared to the Property indicator, the Fight indicator has a smaller threshold and a smaller direct effect slope and therefore a logit closer to zero with a probability closer to 0.5 where the probability curve is steeper, resulting in a larger probability difference for Fight than Property - the direct effect slope size alone does not tell the whole story

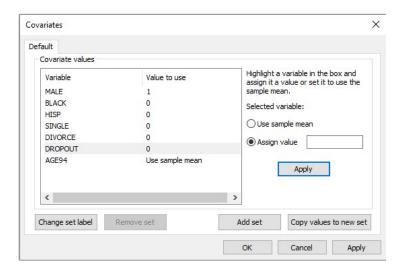
## Mplus Calculator: Computing Direct Effect Probabilities



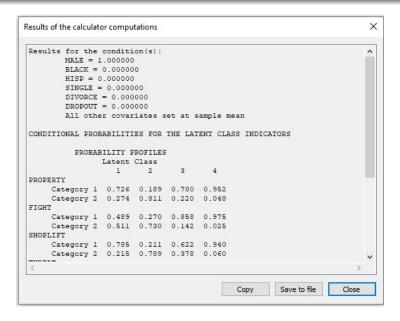
### Calculator Options



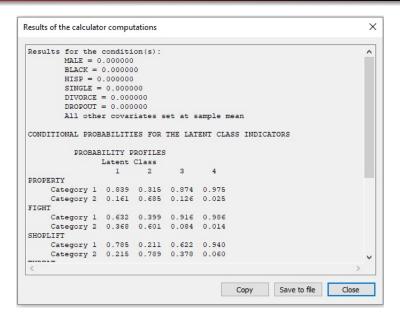
### Calculator Settings: Male=1



#### Calculator Results: Male=1



#### Calculator Results: Male=0



# Calculator Convenience Feature: Copying Sets of Values for Covariates

- Enter your values for the first set
- Click on the button above the OK/Cancel/Apply called "Copy values to new set"
  - That will copy all the values you just entered to a new set
- A second "Default" tab will open to the right of the original Default tab with the copied values that you can edit
  - For clarity, you can change the label of either "Default" tab
- Click OK
- Probabilities for both sets will be shown, one below the other

# Using MODEL CONSTRAINTS for Direct Effect Probability Calculations for the Property Indicator: Class 1 with Male = 0 vs 1, All Other X's=0, Except Age94=Sample Mean

```
MODEL: %OVERALL%
         c ON male-age94:
         property ON male (b1)
         black divorce
                                     MODEL CONSTRAINT:
                                                                NEW(logit0 logit1 prob0 prob1);
         age94 (b2):
                                                                ! suffix of 0/1 corresponds
         fight ON male hisp age94:
                                                                ! to male = 0/1
         shoplift ON black hisp:
                                                                ! sample mean of age94 = 2.957
         threat ON black:
                                                                logit0 = -t1 + b2*2.957:
         pot ON age94;
                                                                logit1 = -t1 + b1 + b2*2.957:
         drug ON black age94:
                                                                prob0 = 1/(1+EXP(-logit0));
         goods ON male black;
                                                                prob1 = 1/(1+EXP(-logit1));
         %c#1%
```

[property\$1] (t1);

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#### Input for PSEM DIFF Priors: Model 10, 15 Direct Effects

TITLE: 7U-7X, 15 direct effects of model 10

PSEM (1.0) with class-varying direct

effects and alignment output

DATA: FILE = asbfree.dat;

VARIABLE: NAMES = property fight shoplift lt50

gt50 force threat injure pot drug

soldpot solddrug con auto bldg goods gambling dsm1-dsm22 sex black hisp single divorce dropout college onset f1

f2 f3 age94 cohort dep abuse;

USEVARIABLES = property fight shoplift threat pot drug goods sex black hisp single divorce dropout age94;

CATEGORICAL = property-goods; CLASSES = c(4);

ANALYSIS: TYPE = MIXTURE:

ESTIMATOR - ML;

STARTS = 400 100; PROCESSORS = 12;

PEEBBORB = 12,

MODEL: %OVERALL%

c ON sex-age94;

property ON sex black divorce age94;

fight ON sex hisp age94; shoplift ON black hisp:

threat ON black;

pot ON age94;

drug ON black age94; goods ON sex black:

%c#1%

property ON sex black divorce age94 (a1-a4);

fight ON sex hisp age94 (a5-a7); shoplift ON black hisp (a8-a9):

threat ON black (a10); pot ON age94 (a11);

drug ON black age94 (a12-a13);

goods ON sex black (a14-a15);

39/58

### Input for PSEM DIFF Priors, Continued

```
%c#2%
property ON sex black divorce age94 (b1-b4):
fight ON sex hisp age94 (b5-b7):
                                                          %c#4%
shoplift ON black hisp (b8-b9):
                                                          property ON sex black divorce age94 (d1-d4):
threat ON black (b10):
                                                          fight ON sex hisp age94 (d5-d7):
pot ON age94 (b11):
                                                          shoplift ON black hisp (d8-d9):
drug ON black age94 (b12-b13):
                                                          threat ON black (d10):
goods ON sex black (b14-b15):
                                                          pot ON age94 (d11):
                                                          drug ON black age94 (d12-d13):
%c#3%
                                                          goods ON sex black (d14-d15):
property ON sex black divorce age94 (c1-c4):
fight ON sex hisp age94 (c5-c7);
                                              MODEL
shoplift ON black hisp (c8-c9);
                                              PRIOR:
                                                          DO(\#,1.15) DIFF(a# b# c# d#)\simALF(0.1.0):
threat ON black (c10):
pot ON age94 (c11):
                                              OUTPUT: ALIGN:
drug ON black age94 (c12-c13):
```

goods ON sex black (c14-c15);

# Results for Class-Varying Direct Effects: Model 10, 15 Direct Effects

- Class-varying direct effects using PSEM 1.0 with DIFF priors: logL = -23,680 (BIC not available)
- Class-varying direct effects using non-PSEM: logL = -23,668, BIC = 48,333
  - 2 fixed direct effects (couldn't be estimated)
- Class-invariant direct effects using non-PSEM: BIC = 48,026
  - Better BIC than for class-varying direct effects
  - No need for class-varying direct effects in this example
- PSEM with DIFF priors useful with class-varying direct effects
  - The ALIGN output option shows which effects have significant differences across classes - see also factor analysis alignment in Asparouhov & Muthén (2014) and other papers at
    - https://www.statmodel.com/ MeasurementInvariance.shtml

# Alignment Output with PSEM DIFF Priors

• a1, b1, c1, d1: Property on sex (male) comparing the 4 classes

DIFF AN	ALYSIS FO	OR PARAMET	ERS			
A1	B1	C1	D1			
Chi-squa	re value	0.053				
Degrees of	of freedom	3				
P-value		0.997				
Param	Param	Value	Value	Difference	SE	P-value
B1	A1	0.723	0.563	0.160	0.750	0.831
C1	A1	0.723	0.563	0.159	0.742	0.830
C1	B1	0.723	0.723	0.000	0.022	0.985
D1	A1	0.724	0.563	0.160	0.754	0.831
D1	B1	0.724	0.723	0.001	0.021	0.980
D1	C1	0.724	0.723	0.001	0.023	0.967
Approxim	nate Invaria	nce Holds For:				
A1	B1	C1	D1			
Average '	Value Acros	s Invariant Para	ameters:	0.683		
Invariant	Values, Dif	ference to Aver	age and S	ignificance		
Param	Value	Difference	SE	P-value		
A1	0.563	-0.120	0.310	0.699		
B1	0.723	0.040	0.458	0.931		
C1	0.723	0.040	0.450	0.930		
D1	0.724	0.040	0.461	0.930		

#### Alignment Output, Continued

• a11, b11, c11, d11: Pot on age94 comparing the 4 classes

DIFF AN	NALYSIS FO	R PARAN	<b>IETERS</b>			
A11	B11	C11	D11			
Chi-squa	re value	7.795				
Degrees	of freedom	3				
P-value		0.050				
Param	Param	Value	Value	Difference	SE	P-value
			,	Difference	OL	1 varac
B11	A11	0.502	0.240	0.261	0.174	0.133
B11 C11	A11 A11	0.502 -0.813	0.240 0.240	Difference		1 /4140
				0.261	0.174	0.133
C11	A11	-0.813	0.240	0.261 -1.053	0.174 0.513	0.133 0.040
C11 C11	A11 B11	-0.813 -0.813	0.240 0.502	0.261 -1.053 -1.315	0.174 0.513 0.555	0.133 0.040 0.018
C11 C11 D11	A11 B11 A11	-0.813 -0.813 0.176	0.240 0.502 0.240	0.261 -1.053 -1.315 -0.064	0.174 0.513 0.555 0.047	0.133 0.040 0.018 0.178

Approximate Invariance Was Not Found.

• c11 vs b11 shows the largest difference: Pot on age94, comparing the third and second classes - Pot and High

#### Outline

- Background
- Antisocial behavior example
- Searching for direct effects
  - Different approaches, Penalized SEM (PSEM)
  - ASB example (reduced version)
    - Latent classes and C ON X results
- Evaluating direct effects
  - Direct effects results, calculator
- Class-varying direct effects
  - PSEM DIFF priors, ALIGN option
- Multistep analysis with direct effects and distal outcomes
- Recap and further research

# Multistep Analysis with Direct Effects and Distal Outcomes

- Adding distal outcomes to the analysis with direct effects
  - The first step measurement model includes the covariates and their direct effects
  - What should the last step look like?
    - 3-step
    - BCH
    - 2-step
- Example: ASB data with the 15 direct effects of Model 10, adding alcohol dependence and abuse as distal outcomes
  - Strongly skewed variables with large floor effects; treated as continuous in these analyses
  - Likely to be correlated even when conditioned on latent class

## First Step, Combined Approach with Direct Effects

TITLE:

First step, combined approach, measure-

ment model including direct effects

DATA: FILE = asbfree.dat;

VARIABLE: NAMES = property fight shoplift lt50

gt50 force threat injure pot drug soldpot solddrug con auto bldg goods gambling dsm1-dsm22 sex black hisp single divorce dropout college onset f1 f2 f3

age94 cohort dep abuse;

USEVAR = property fight shoplift threat pot drug goods sex black hisp single

divorce dropout age94;

AUXILIARY = dep abuse;

CATEGORICAL = property-goods;

CLASSES = c(4);

ANALYSIS: TYPE = MIXTURE;

ESTIMATOR = ML; OPTSEED = 21345;

PROCESSORS = 12;

MODEL: %OVERALL%

c ON sex-age94:

! Direct effects:

property ON sex black divorce age94;

fight ON sex hisp age94; shoplift ON black hisp; threat ON black:

pot ON age94; drug ON black age94;

goods ON sex black;

**SAVEDATA: SAVE = cprob bchweights**;

FILE = final.dat;

OUTPUT: ! Re-ordering the classes

! based on previous run: SVALUES(1 4 2 3);

3 VALUES(1 4 2 3),

#### Input for 3-Step: Last Step

ANALYSIS: TYPE = MIXTURE; ESTIMATOR = ML:

TITLE:

Last step of 3-step

Last step of 3-step

Last step of 3-step

DATA: FILE = final dat:

VARIABLE: NAMES = property fight shoplift threat

MODEL: %OVERALL%

pot drug goods sex black hisp single divorce dropout age94 dep abuse w1-w4

dep abuse ON sex-age94;

cprob1-cprob4 n: dep WITH abuse;

! No direct effects since indicators

USEVARIABLES = sex-age94 ! not in the model dep abuse n:

NOMINAL [n#1@1.616 n#2@-1.634 n#3@-1.108];

NOMINAL = n; [11#1@1.010 11#2@-1.034 11#3@-1.108]; %c#2%

CLASSES = c(4); [n#1@4.971 n#2@6.760 n#3@4.382]; %c#3%

[n#1@-0.101 n#2@-0.734 n#3@1.908];

%c#4%

 $[n\#1@-2.518\ n\#2@-9.861\ n\#3@-3.323];$ 

%c#1%

### Input for BCH: Last Step

TITLE:

Last step of BCH

DATA: FILE =final.dat; ANALYSIS: TYPE = MIXTURE;

VARIABLE: NAMES = property fight shoplift threat

pot drug goods sex black hisp single divorce dropout age94 dep abuse w1-w4

cprob1-cprob4 n:

USEVARIABLES = dep abuse

sex-age94 w1-w4:

TRAINING = w1-w4(BCH);

CLASSES = c(4);

ESTIMATOR = ML:

STARTS = 0:

DIANTS = 0,

PROCESSORS = 12;

MODEL: %OVERALL%

dep abuse ON sex-age94;
dep WITH abuse;

! No direct effects since indicators

! not in the model

#### Input for 2-Step: Last Step

TITLE:

Last step of 2-step

DATA:

FILE IS asbfree.dat;

VARIABLE:

NAMES = property fight shoplift lt50 gt50 force threat injure pot drug

soldpot solddrug con auto bldg goods gambling dsm1-dsm22 sex black hisp single divorce dropout college onset f1

f2 f3 age94 cohort dep abuse;

USEVARIABLES = property fight shoplift threat pot drug goods sex black hisp single divorce dropout age94 dep abuse:

CATEGORICAL = property-goods;

CLASSES = c(4);

ANALYSIS:

 $\mathsf{TYPE} = \mathsf{MIXTURE};$ 

ESTIMATOR - ML;

STARTS = 0; ! STARTS=400 100 gives

another solution

PROCESSORS = 12;

MODEL:

%OVERALL%

dep abuse ON sex-age94;

dep WITH abuse;

! Input continues on next slides

#### Input for 2-Step, Last Step, Continued

```
! Direct effects need to be included
                                                       c#1 ON sex@1.28575:
! because the indicators are in the model
                                                       c#1 ON black@1.26114:
                                                       c#1 ON hisp@-0.25585:
! Output from SVALUES.
                                                       c#1 ON single@0.01131;
! First for OVERALL.
                                                       c#1 ON divorce@0.27248:
! then class-specific
                                                       c#1 ON dropout@0.28195;
property ON sex;
                                                       c#1 ON age94@-0.20425;
                                                       c#2 ON sex@1.91380:
property ON black;
property ON divorce;
                                                       c#2 ON black@0.52702:
property ON age94;
                                                       c#2 ON hisp@-0.65689;
                                                       c#2 ON single@0.43056;
fight ON sex;
                                                       c#2 ON divorce@0.79212:
fight ON hisp;
fight ON age94;
                                                       c#2 ON dropout@0.62244;
shoplift ON black;
                                                       c#2 ON age94@-0.23210;
shoplift ON hisp;
                                                       c#3 ON sex@-0.01762:
threat ON black:
                                                       c#3 ON black@-1.07972:
                                                       c#3 ON hisp@-0.66859;
pot ON age94;
                                                       c#3 ON single@0.31222;
drug ON black;
drug ON age94;
                                                       c#3 ON divorce@0.46191:
goods ON sex;
                                                       c#3 ON dropout@-0.06458;
goods ON black;
                                                       c#3 ON age94@-0.03458;
```

#### Input for 2-Step Continued

```
[ c#1@-0.64931 ]:
[ c#2@-1.86032 ]:
[ c#3@-0.49162 ]:
%C#1%
property ON sex@0.67567 (29):
property ON black@-0.74400 (30):
property ON divorce@-0.45823 (31):
property ON age94@-0.09999 (32):
fight ON sex@0.58659 (33):
fight ON hisp@0.73120 (34):
fight ON age94@-0.10968 (35):
shoplift ON black@-0.16065 (36);
shoplift ON hisp@0.45153 (37);
threat ON black@-0.71839 (38);
pot ON age94@0.22548 (39);
drug ON black@-0.81339 (40);
drug ON age94@0.22387 (41);
goods ON sex@0.83449 (42);
goods ON black@-0.54608 (43);
```

```
[ property$1@1.35209 ];
[ fight$1@0.21694 ];
[ shoplift$1@1.29552 ];
[ threat$1@-0.69313 ];
[ pot$1@1.35039 ];
[ drug$1@3.95855 ];
[ goods$1@2.79884 ];
```

#### Input for 2-Step Continued

#### %C#2%

```
property ON sex@0.67567 (29):
property ON black@-0.74400 (30);
property ON divorce@-0.45823 (31):
property ON age94@-0.09999 (32):
fight ON sex@0.58659 (33):
fight ON hisp@0.73120 (34):
fight ON age94@-0.10968 (35):
shoplift ON black@-0.16065 (36);
shoplift ON hisp@0.45153 (37);
threat ON black@-0.71839 (38):
pot ON age94@0.22548 (39):
drug ON black@-0.81339 (40):
drug ON age94@0.22387 (41):
goods ON sex@0.83449 (42):
goods ON black@-0.54608 (43):
```

```
[ property$1@-1.07329 ];
 [ fight$1@-0.73433 ];
 [ shoplift$1@-1.31987 ];
 [ threat$1@-1.69122 ];
 [ pot$1@-1.69565 ];
 [ drug$1@0.13231 ];
 [ goods$1@0.34741 ];
```

# Comparing Results for 3-Step, BCH, and 2-Step

- Class probabilities from the first step (Model 10, 15 direct effects):
  - 0.321 (Threat/Fight), 0.130 (High), 0.172 (Drugs), 0.377 (low)
- Class probabilities in last step (STARTS=0):
  - 3-step: 0.461, 0.019, 0.164, 0.356
    - Failure simulations in Asparouhov & Muthén (2014; Web Note 15) show undesirable influence of continuous distal outcomes
  - BCH: 0.321, 0.130, 0.173, 0.377
  - 2-step: 0.321, 0.130, 0.172, 0.377
- STARTS > 0
  - 3-step: Better LL but not replicated, class probabilities still different from first step, different estimates for distal outcomes regressed on covariates - solution influenced by distal outcomes
  - BCH: Same solution as for STARTS=0
  - 2-step: Better LL, same class probabilities, different estimates for distal outcomes regressed on covariates - solution influenced by distal outcomes
- BCH preferrable (see also Asparouhov & Muthén, Web Note 21 and Web Talk 8) due to within-class non-normality of distal outcomes

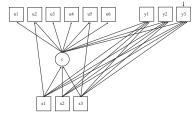
# Distal Outcomes for BCH: Comparing Y ON X with Direct Effects in 1st Step vs No Covariates, No Directs in 1st Step

1st step: Covariates and 15 direct effects (Model 10)				1st step: Latent class indicators only		
	Estimate	S.E.	Est./S.E.	Estimate	S.E.	Est./S.E.
DEP ON						
SEX	0.140	0.027	5.147	0.158	0.025	6.248
BLACK	-0.031	0.032	-0.979	0.001	0.028	0.018
HISP	-0.008	0.032	-0.255	-0.022	0.032	-0.702
SINGLE	0.250	0.029	8.585	0.251	0.029	8.647
DIVORCE	0.154	0.032	4.863	0.163	0.032	5.171
DROPOUT	0.362	0.036	10.171	0.374	0.035	10.558
AGE94	0.004	0.006	0.801	0.000	0.006	0.031
ABUSE ON						
SEX	0.268	0.032	8.386	0.274	0.030	9.264
BLACK	-0.294	0.037	-7.939	-0.269	0.033	-8.105
HISP	-0.051	0.038	-1.363	-0.094	0.037	-2.529
SINGLE	0.353	0.034	10.310	0.359	0.034	10.525
DIVORCE	0.257	0.037	6.913	0.277	0.037	7.493
DROPOUT	0.150	0.042	3.591	0.169	0.042	4.060
AGE94	-0.013	0.007	-1.909	-0.027	0.007	-4.092

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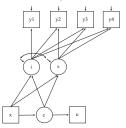
### Recap of Alternative Approaches



- Analyses without the distal outcomes
  - (1) Including direct effects
  - (2) Ignoring direct effects, including X's
  - Approaches (1) and (2) were compared with respect to C ON X
- Analyses with the distal outcomes: Multistep such as BCH
  - (3) Including direct effects in first step
  - (4) Ignoring direct effects in first step
  - (5) Including direct effects only in last step (for 2-step)
  - (6) Excluding covariates in first step (measurement model using indicators only), adding them in last step, ignoring direct effects
  - Approaches (3) and (6) were compared with respect to Y ON X

#### **Further Research**

- Monte Carlo simulation studies
  - PSEM approach
  - Figure 12 of Asparouhov & Muthén (2025)
  - MplusAutomation https: //www.statmodel.com/usingmplusviar.shtml
- PSEM needs work for the ordinal case (avoiding threshold collapse)
- Other measurement models:
  - Growth mixture modeling:
    - Muthén & Shedden (1999), Muthén et al. (2002), Muthén (2004),
       Asparouhov & Muthén (2014; Web Note 15)



#### References

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